

Joint Gaussian Distribution



Gaussians everywhere are due to CLT. Lots of little effects add up.

Holds not just for single random variables but for two or more correlated random variables taken as an iid sequences.

$(U_1, V_1), (U_2, V_2), \dots$ For example height/weight of elephants seen on safari.

Suppose each is mean-zero for convenience.

Then as $n \rightarrow \infty$, $\left(\frac{U_1 + \dots + U_n}{\sqrt{n}}, \frac{V_1 + \dots + V_n}{\sqrt{n}}\right)$ has limiting bivariate distribution. By CLT this will be jointly Gaussian (jG).

Def: Random variables X and Y are jG if every linear combination $aX + bY$ is a Gaussian r.v.

$$f_{X,Y}(u,v) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left\{-\frac{\left(\frac{u-\mu_x}{\sigma_x}\right)^2 + \left(\frac{v-\mu_y}{\sigma_y}\right)^2 - 2\rho\left(\frac{u-\mu_x}{\sigma_x}\right)\left(\frac{v-\mu_y}{\sigma_y}\right)}{2(1-\rho^2)}\right\}$$

Simplest setting is independent standard normal r.v.: W, Z

$$f_{W,Z}(\alpha, \beta) = \left(\frac{e^{-\frac{\alpha^2}{2}}}{\sqrt{2\pi}}\right) \left(\frac{e^{-\frac{\beta^2}{2}}}{\sqrt{2\pi}}\right) = \frac{e^{-\frac{\alpha^2 + \beta^2}{2}}}{2\pi}$$

where $\mu_W = \mu_Z = 0$, $\sigma_x = \sigma_y = 1$, $\rho = 0$.

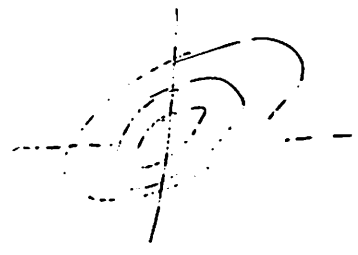
To get back general result:

$$\begin{bmatrix} X \\ Y \end{bmatrix} = A \begin{bmatrix} W \\ Z \end{bmatrix} + \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}$$



where $A = \begin{bmatrix} \sqrt{\frac{\sigma_x^2(1+\rho)}{2}} & -\sqrt{\frac{\sigma_y^2(1-\rho)}{2}} \\ \sqrt{\frac{\sigma_y^2(1+\rho)}{2}} & \sqrt{\frac{\sigma_x^2(1-\rho)}{2}} \end{bmatrix}$.

level sets are ellipses



Key properties

Suppose X, Y bivariate normal with $\mu_x, \mu_y, \sigma_x, \sigma_y, \rho$.

- (a) $X \sim N(\mu_x, \sigma_x^2)$ and $Y \sim N(\mu_y, \sigma_y^2)$.
- (b) any linear combination, $aX + bY$ is Gaussian r.v., i.e. X and Y are j.c.
- (c) ρ is correlation coeff between X and Y .
- (d) $X \perp\!\!\!\perp Y$ iff $\rho = 0$.
- (e) For estimation of Y from X , $L^*(X) = g^*(x)$, which is to any, linear estimators are good enough in Gaussian case
- (f) Conditional distribution of Y given $X = u$ is $N(L^*(u), \sigma_c^2)$
 where σ_c^2 is MSE for LMSE estimator

To prove, restrict to $\mu_x = \mu_y = 0, \sigma_x = \sigma_y = 1$.

Then $f_{XY}(u, v) = \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left\{-\frac{u^2 + v^2 - 2\rho uv}{2(1-\rho^2)}\right\}$

$$= \left[\frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u^2}{2}\right) \right] \left[\frac{1}{\sqrt{2\pi(1-\rho^2)}} \exp\left(-\frac{(v-\rho u)^2}{2(1-\rho^2)}\right) \right]$$

function of u alone
function of v from fixed u is $N(\rho u, 1-\rho^2)$.

In particular first factor is $f_X(u)$ and second factor is $f_{Y|X}(v|u)$.

Since given $X=u$, conditional distribution of Y is $N(\mu, 1-p^2)$.

$$g^*(u) = E[Y|X=u] = \mu$$

We can separately compute $L^*(u)$ and see it is also μ , so ~~therefore~~

LMSE coincides with MMSE estimator.

In higher dimensions

$$f_X(\vec{u}) = \frac{1}{(2\pi)^{n/2} \sqrt{K + \Sigma}} \exp\left(-\frac{(\vec{u}-\mu)^T \Sigma^{-1} (\vec{u}-\mu)}{2}\right).$$