

Laws of Large Numbers

(concentration of measure = dust.)

→ Statements about sample mean of many random variables, and how it converges to true mean, $S_n = X_1 + X_2 + \dots + X_n$.

Prop: Suppose X_1, X_2, \dots is sequence of uncorrelated random variables s.t. each X_k has finite mean μ and variance less than or equal to C .

Then for any $\delta > 0$,

$$\Pr \left[\left| \frac{S_n}{n} - \mu \right| \geq \delta \right] \leq \frac{C}{n\delta^2} \quad (\text{convergence in probability})$$

and so this $\rightarrow 0$ as $n \rightarrow \infty$.

Proof: The mean of S_n/n is

$$E \left[\frac{S_n}{n} \right] = E \left[\frac{\sum_{k=1}^n X_k}{n} \right] = \frac{\sum_{k=1}^n E[X_k]}{n} = \frac{n\mu}{n} = \mu.$$

The variance can be as follows:

$$\text{var} \left(\frac{S_n}{n} \right) = \text{var} \left(\frac{\sum_{k=1}^n X_k}{n} \right) = \frac{\sum_{k=1}^n \text{var}(X_k)}{n^2} \leq \frac{nC}{n^2} = \frac{C}{n}.$$

now recall Chebyshev's ineq. if Z is random variable with finite mean μ and variance σ^2 , then for any $d > 0$, $\Pr [|X - \mu| \geq d] \leq \frac{\sigma^2}{d^2}$.

so apply to $Z = \frac{S_n}{n}$, we get:

$$\Pr \left[\left| \frac{S_n}{n} - \mu \right| \geq \delta \right] \leq \frac{C}{n\delta^2}.$$

⇒ matlab simulation for $\frac{S_n}{n}$ vs. n and S_n vs. n for $U(0,1)$ iid.

hard draw, Z with radio telephony experiments

* Suppose X_1, \dots, X_{100} are each r.v. with $\mu = 5$ and $\sigma^2 = 1$. Also, $|\text{cov}(X_i, X_j)| \leq 0.1$ if $i = j \pm 1$ and $\text{cov}(X_i, X_j) = 0$ if $|i - j| \geq 2$.

Let $S_{100} = X_1 + \dots + X_{100}$

(a) show $\text{var}(S_{100}) \leq 120$.

(b) find upper bound on $\text{Pr}\left[\left|\frac{S_{100}}{100} - 5\right| \geq 0.5\right]$.

(a) for any n , we know $\text{var}(S_n) = \sum_{i=1}^n \sum_{j=1}^n \text{cov}(X_i, X_j)$.

$$= \sum_{i=1}^{n-1} \text{cov}(X_i, X_{i+1}) + \sum_{i=1}^n \text{cov}(X_i, X_i) + \sum_{i=2}^n \text{cov}(X_i, X_{i-1})$$

$$\leq (n-1)(0.1) + n + (n-1)(0.1) < 1.2n.$$

(b) from part (a), get that $\text{var}\left(\frac{S_{100}}{100}\right) = \frac{1}{100^2} \text{var}(S_{100}) \leq 0.012$.

Also know $E\left[\frac{S_{100}}{100}\right] = \mu = 5$. So applying Chebyshev:

$$\text{Pr}\left[\left|\frac{S_{100}}{100} - 5\right| \geq 0.5\right] \leq \frac{0.012}{(0.5)^2} = 0.048.$$

* Let U_1, \dots, U_n be iid $E(\lambda)$ with unknown λ . (has mean $\frac{1}{\lambda}$ and variance $\frac{1}{\lambda^2}$).

The ML estimator turns out to be $\hat{\lambda}_{ML} = \frac{n}{S_n}$, where $S_n = U_1 + \dots + U_n$.

Find the number of observations n s.t. $[(0.9)\hat{\lambda}_{ML}, (1.1)\hat{\lambda}_{ML}]$ is confidence interval for estimating λ with confidence level 96%.

Want $\text{Pr}\left[\frac{0.9n}{S_n} \leq \lambda \leq \frac{1.1n}{S_n}\right] > 0.96$ or equivalently $\text{Pr}\left[0.9 \leq \frac{\lambda S_n}{n} \leq 1.1\right] \geq 0.96$.

To use Chebyshev, note that $E\left[\frac{\lambda S_n}{n}\right] = E[\lambda U_1] = 1$.

$$\text{var}\left[\frac{\lambda S_n}{n}\right] = \frac{\lambda^2}{n^2} \text{var}(S_n) = \frac{\lambda^2 \text{var}(U_1)}{n} = \frac{1}{n}.$$

So $\text{Pr}\left[\left|\frac{\lambda S_n}{n} - 1\right| \geq \delta\right] \leq \frac{1}{n\delta^2}$. Setting $\frac{1}{n\delta^2} = 0.04$ with $\delta = 0.1$ gives $n = \frac{1}{(0.1)^2(0.04)} = 2500$.

Let X_1, X_2, \dots, X_n be independent random variables with $E(X_i) = \mu$ and $\text{Var}(X_i) = \sigma^2$.

$$S_n = \sum_{i=1}^n X_i \quad \text{and} \quad \bar{X}_n = \frac{S_n}{n}$$

Then $E(S_n) = n\mu$ and $\text{Var}(S_n) = n\sigma^2$.

$$E(\bar{X}_n) = \mu \quad \text{and} \quad \text{Var}(\bar{X}_n) = \frac{\sigma^2}{n}$$

$$P(\bar{X}_n - \mu \leq -\epsilon) = P(S_n - n\mu \leq -n\epsilon)$$

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By Chebyshev's inequality, $P(|\bar{X}_n - \mu| \geq \epsilon) \leq \frac{\text{Var}(\bar{X}_n)}{\epsilon^2} = \frac{\sigma^2}{n\epsilon^2}$.

$$P(|\bar{X}_n - \mu| < \epsilon) \geq 1 - \frac{\sigma^2}{n\epsilon^2}$$

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As $n \rightarrow \infty$, $P(|\bar{X}_n - \mu| < \epsilon) \rightarrow 1$.

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We know convergence of sample mean to mean using LLN; what about full distributional form of the sum?

What did De Moivre-Laplace then say?

Central limit theorem. Suppose X_1, X_2, \dots are iid r.v. with mean μ and variance σ^2 . Let $S_n = X_1 + \dots + X_n$. Then:

$$\lim_{n \rightarrow \infty} \Pr \left[\frac{S_n - n\mu}{\sqrt{n\sigma^2}} \leq c \right] = \Phi(c).$$

~~Ex~~. Gaussian approx.

Ex. Suppose each day of the year, a stock:

}	increase by 1%	w.p. 0.5
	stays	w.p. 0.4
	decrease by 1%	w.p. 0.1.

Consider value of 1 stock after 365 days.

- (a) Probability stock at least triples.
- (b) Probability stock at least quadruples.
- (c) median value after one year.

Value after a year $Y = D_1 D_2 \dots D_{365}$ where D_k is growth factor for day

k . Convert to log to use CLT:

$$\ln Y = \sum_{k=1}^{365} \ln D_k.$$

$$\mu = E[\ln(D_k)] = 0.5 \ln 1.01 + 0.4 \ln 1 + 0.1 \ln 0.99 = 0.00397.$$

$$\sigma^2 = \text{var}(\ln(D_k)) = E[(\ln D_k)^2] - \mu^2 = 0.00004384$$

$\ln Y$ is approx Gaussian with mean $365\mu = 1.450$
~~var~~ std $\sqrt{365\sigma^2} = 0.127$

So $P[Y \geq c] = Pr[\ln(Y) \geq \ln(c)]$
 $= Pr\left[\frac{\ln(Y) - 1.450}{0.127} \geq \frac{\ln(c) - 1.450}{0.127}\right]$
 $\approx Q\left(\frac{\ln(c) - 1.450}{0.127}\right)$

(a) In particular $Pr(Y \geq 3) = Q(-2.77) = 0.997$
 (b) $Pr(Y \geq 4) = Q(-0.4965) = 0.69$

(c) median value is value of c s.t. $Pr(Y \geq c) = 0.5$
 which by Gaussian approx is $e^{1.450} = 4.26$

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$$f(x) = \frac{1 - (x-1)^2}{1 + (x-1)^2}$$

$$f(x) = \frac{2 - (x-1)^2}{1 + (x-1)^2}$$

(a) In question $f(x) = \frac{1 - (x-1)^2}{1 + (x-1)^2} = 0.999$

(b) $f(x) = \frac{1 - (x-1)^2}{1 + (x-1)^2} = 0.999$

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where $f(x) = \frac{1 - (x-1)^2}{1 + (x-1)^2}$