

## Sums of random variables

We know by linearity,  $E[X+Y] = E[X] + E[Y]$  but what about the full distributional form of  $X+Y$ ?

This problem often arises when considering additive measurement noise:

$$S \xrightarrow{\oplus} Z \quad Z = S + N, \text{ Signal plus noise.}$$

$\downarrow$

Sum of two dice

Sums of integer-valued random variables:

Let  $S = X+Y$ , want  $p_S$  in terms of  $p_X$  and  $p_Y$ .

For fixed value  $k$ , possible ways to get  $S=k$  can be indexed according to the value of  $X$ . For  $S=k$  to happen, have to have  $X=j$  and  $Y=k-j$ , for some  $j$ .

By total probability,

$$\begin{aligned} p_S(k) &= P[X+Y=k] \\ &= \sum_j P[X=j, Y=k-j] \\ &= \sum_j p_{XY}(j, k-j). \end{aligned}$$

If  $X \perp\!\!\!\perp Y$ , then  $p_{XY}(j, k-j) = p_X(j)p_Y(k-j)$  and so

$$p_S(k) = \sum_j p_X(j)p_Y(k-j) \quad \text{which is the convolution operation.}$$

$$p_S(k) = p_X * p_Y.$$

Example: Let  $X \perp\!\!\!\perp Y$ , with

$$p_X(x) = \begin{cases} 1/3, & x=1,2,3 \\ 0, & \text{else} \end{cases}$$

$$p_Y(y) = \begin{cases} 1/2, & y=0 \\ 1/3, & y=1 \\ 1/6, & y=2 \\ 0, & \text{else} \end{cases}$$

To calculate pmf of  $S = X+Y$ , note that range of possible values of  $S$  are integers  $1, 2, \dots, 5$ , so  $p_S(s) = 0$  for  $s \notin 1, 2, \dots, 5$ .

For other values, let us calculate, one-by-one:

$$p_S(1) = \sum_x p_X(x) p_Y(1-x) = p_X(1)p_Y(0) = \frac{1}{3} \cdot \frac{1}{2} = \frac{1}{6}.$$

$$p_S(2) = p_X(1)p_Y(1) + p_X(2)p_Y(0) = \frac{1}{3} \cdot \frac{1}{2} + \frac{1}{3} \cdot \frac{1}{2} = \frac{5}{18}$$

$$p_S(3) = p_X(1)p_Y(2) + p_X(2)p_Y(1) + p_X(3)p_Y(0) = \frac{1}{3} \cdot \frac{1}{6} + \frac{1}{3} \cdot \frac{1}{3} + \frac{1}{3} \cdot \frac{1}{2} = \frac{1}{3}$$

$$p_S(4) = p_X(2)p_Y(2) + p_X(3)p_Y(1) = \frac{1}{3} \cdot \frac{1}{6} + \frac{1}{3} \cdot \frac{1}{3} = \frac{1}{6}$$

$$p_S(5) = p_X(3)p_Y(2) = \frac{1}{3} \cdot \frac{1}{6} = \frac{1}{18}.$$

Suppose  $X$  has binomial distribution with parameters  $n$  and  $p$ .  
 $\quad\quad\quad Y \qquad\qquad n \qquad p$

$X \perp\!\!\!\perp Y$ , find distribution of  $X+Y$ .

Recall  $X$  is #heads in  $n$  tosses,  $Y$  is #heads in  $n$  tosses,

Thus  $S = X+Y$  is #heads in  $n+n$  tosses, i.e. binomial with  $(n+n, p)$ .

$X \sim \text{Poisson}(\lambda_1)$   $X \perp\!\!\!\perp Y$ , find distribution of  $S = X+Y$   
 $\quad\quad\quad Y \sim \text{Poisson}(\lambda_2)$

Since Poisson  $\sim$  binomial, expect should be Poisson with  $\lambda_1 + \lambda_2$ .

Do formally using convolution:

$$p_S(k) = \sum_{j=0}^k \left( \frac{\lambda_1^j e^{-\lambda_1}}{j!} \right) \left( \frac{\lambda_2^{k-j} e^{-\lambda_2}}{(k-j)!} \right) = \left( \sum_{j=0}^k \frac{\lambda_1^j \lambda_2^{k-j}}{j! (k-j)!} \right) e^{-(\lambda_1 + \lambda_2)}$$

$$= \left( \sum_{j=0}^k \binom{k}{j} p^j (1-p)^{k-j} \right) \frac{\lambda^{k-p}}{k!} \quad \text{where } \lambda = \lambda_1 + \lambda_2 \text{ and } p = \frac{\lambda_1}{\lambda}$$

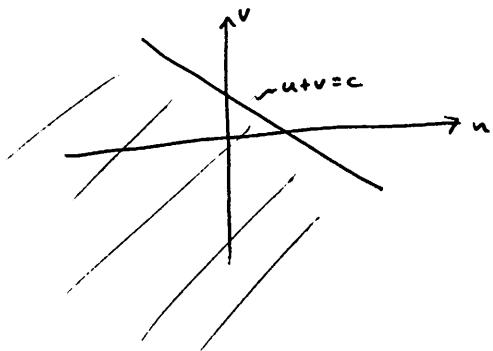
$$= \frac{\lambda^k e^{-\lambda}}{k!} \quad \text{since } \sum_{j=0}^k \binom{k}{j} p^j (1-p)^{k-j} \text{ is binomial sum.}$$

Continuous random variables.

$S = X+Y$ , want to express  $f_S$  in terms of  $f_{XY}$ .

- First find cdf of  $S$ , then differentiate to get pdf.

- For any  $c \in \mathbb{R}$ ,  $\{S \leq c\}$  is same as point  $(x,y)$  falls in lower diagonal:



Integrating over a region is by picking one dimension, then next:

$$F_S(c) = \Pr[S \leq c] = \int_{-\infty}^{\infty} \int_{-\infty}^{c-u} f_{XY}(u,v) dv du$$

taking derivative:

$$\begin{aligned} f_S(c) &= \frac{dF_S(c)}{dc} = \int_{-\infty}^{\infty} \frac{d}{dc} \left( \int_{-\infty}^{c-u} f_{XY}(u,v) dv \right) du \\ &= \int_{-\infty}^{\infty} f_{XY}(u, c-u) du. \end{aligned}$$

Now since  $X \perp\!\!\!\perp Y$ , recover convolution:

$$= \int_{-\infty}^{\infty} f_X(u) f_Y(c-u) du$$

$$f_S = f_X * f_Y.$$

Consider  $X \sim N(0, \sigma^2)$ ,  $Y \sim N(0, \sigma^2)$ ,  $X \perp\!\!\!\perp Y$

find pdf of  $S = X+Y$ .

$$\begin{aligned}
 f_S(s) &= \int_{-\infty}^{\infty} f_X(u) f_Y(c-u) du \\
 &= \int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{u^2}{2\sigma^2}} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(c-u)^2}{2\sigma^2}} du \\
 &= \frac{1}{2\pi\sigma^2} \int_{-\infty}^{\infty} e^{-\frac{u^2}{2\sigma^2} - \frac{(c-u)^2}{2\sigma^2}} du \\
 &= \frac{1}{2\pi\sigma^2} \int_{-\infty}^{\infty} \exp\left\{\frac{1}{2\sigma^2}(-u^2 - c^2 + 2cu - u^2)\right\} du \\
 &= \frac{1}{2\pi\sigma^2} \int_{-\infty}^{\infty} \exp\left\{\frac{1}{2\sigma^2}( )\right\} \exp\left\{\frac{1}{2\sigma^2}( )\right\} du.
 \end{aligned}$$

Want pieces such that they go away, i.e. integrate to 1, since known Gaussian pdfs.

To do so, use "Completing the Square": see notes.

eventually get

$$f_S(s) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{s^2}{2\sigma^2}} \quad \text{when } \sigma^2 = 2\sigma^2.$$

Sum of two Gaussians is also Gaussian, with variance as sum of variances.

$\alpha$ -stable distributions are not so common.