

example

Time until small meteorite first lands in Sahara desert is modeled as exponential r.v. with mean 10 days. It is currently midnight. What is probability that a meteorite first lands some time between 6am and 6pm of first day?

Let X be time until event, measured in days.

Then X is $\sim \exp(\lambda = \frac{1}{10})$ since $\text{mean} = \frac{1}{\lambda} = 10$.

$$\text{Want } P[X_1 \leq x \leq \frac{3}{4}] = P[x \leq \frac{1}{4}] - P[x > \frac{1}{4}] = e^{-\frac{1}{4}} - e^{-\frac{3}{4}} = 0.0476.$$

where we used the survival function.

example Let T be an exponentially distributed r.v. with $\lambda = \ln 2$. Find an expression for $P(T \geq t)$ as a function of t hr, $t \geq 0$, and find $P(T \leq 1 | T \leq 2)$

$$P(T \geq t) = F_T^c(t) = e^{-\lambda t} = e^{-(\ln 2)t} = 2^{-t}$$

$$P(T \leq 1 | T \leq 2) = \frac{P(T \leq 1, T \leq 2)}{P(T \leq 2)} = \frac{P(T \leq 1)}{P(T \leq 2)} = \frac{1 - \frac{1}{2}}{1 - \frac{1}{4}} = \frac{2}{3}$$

Bayes

Memoryless property of exponential

Suppose T is exponentially distributed with parameter λ , then $P(T>t) = e^{-\lambda t}$.

$$\text{Thus: } P(T>s+t | T>s) = \frac{P(T>s+t, T>s)}{P(T>s)} = \frac{P(T>s+t)}{P(T>s)} = \frac{e^{-\lambda(s+t)}}{e^{-\lambda s}} = e^{-\lambda t} \cdot P(T>s)$$

$\therefore P(T>s+t | T>s) = P(T>t)$. This is the memoryless property.

Interpretation in terms of lifetime: having not failed up to time s doesn't affect future time.

In fact exponential is one of the only nontrivial distributions to have this property.

$P(T>s+t | T>s) = P(T>t)$ can be rewritten as

$$P(T>s+t) = P(T>s) P(T>t) \quad \text{by using fact } (T>s+t) \cap T(s) = (T>s+t).$$

Let $g(x) = \log P(T>x)$ as is legitimate for positive $P(T>x)$, we get:

$$g(s+t) = g(s) + g(t), \quad \text{which is the Cauchy equation.}$$

one solution is $g(x) = \beta x$. (There are highly pathological other solutions that are unbounded in any finite interval and not measurable).

$$\text{If } g(x) = \beta x, \text{ then } P(T>x) = e^{\beta x}$$

which yields $T \sim \mathcal{E}(\beta)$.

Exponential distribution is continuous-time analogue of geometric distribution.

Similarly we can go from Bernoulli Process to Poisson Process.

Conversion is through a time-scaling operation.

\Rightarrow see notes.

	Poisson	Bernoulli
Times of Arrival	continuous	discrete
pmf of # arrivals	Poisson	Binomial
interarrival cdf	Exponential	Geometric
arrival rate	λ / unit time	p / per trial

exaplh 2 You get email according to a Poisson process at a rate of $\lambda = 0.2$ messages/hour. You check email every hour. What is probability of finding 0 and 1 new messages?

Use Poisson pmf: $\frac{e^{-\lambda \tau} (\lambda \tau)^k}{k!}$ with $\tau = 1$, $k=0$ or $k=1$.

$$P(0,1) = e^{-0.2} \cdot 0.819 \quad P(1,1) = 0.2 e^{-0.2} \cdot 0.164$$

Suppose you have not checked email for a whole day: what is probability of no new messages?

$$P(0,24) = e^{-0.2 \cdot 24} = 0.0083.$$

exaplh. 1

Consider a Poisson process on $[0, T]$ with rate $\lambda > 0$, let $0 < t < T$.

Defn X_1 to be number of counts during $[0, t]$, X_2 to be # counts during $[t, T]$, X to be total counts during $[0, T]$.

Let i, j, n be nonnegative integers s.t. $n = i + j$.

Find $\Pr[X=n]$ in terms of $n, i, j, \tau, T, \lambda$:

$$\frac{e^{-\lambda \tau} (\lambda \tau)^n}{n!}$$

$$\text{Find } \Pr[X_1=i] = \frac{e^{-\lambda \tau} (\lambda \tau)^i}{i!}$$

$$\text{Find } \Pr[X_2=j] = \frac{e^{-\lambda(T-\tau)} (\lambda(T-\tau))^j}{j!}$$